

# Modeling Age-Specific Mortality for Countries with Generalized HIV Epidemics\*

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## Abstract

Population projections that forecast the future size and age-composition of a country are crucial tools for appropriately planning the future allocation of societal resources. A projection model for countries with generalized HIV epidemics should take into account the future trajectory of the epidemic given the severe effect a generalized epidemic can have on the mortality conditions and composition of a population. We present a model of age-specific mortality as a function of HIV prevalence and one of three possible inputs (life expectancy at birth, child mortality alone, or child mortality with adult mortality) for the 39 countries of the world experiencing a generalized HIV epidemic. We perform an in-sample validation where results show slight errors for several mortality indicators. Combined with the outputs of existing epidemiological and demographic models, this model makes it possible to estimate future mortality profiles for countries with generalized HIV epidemics.

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## Motivation

Population projections that forecast the future size and age-composition of a country or region are crucial tools for appropriately planning the future allocation of societal resources such as medical, educational, or government services. These kinds of projections for countries with generalized HIV epidemics ( $> 1\%$  prevalence in the general population and no concentration of the epidemic in high-risk subgroups such as IV drug users, men who have sex with men, or sex workers) are especially useful given the dramatic effect a generalized epidemic can have on life expectancy and the age-composition of a population (Ngom and Clark, 2003; Blacker, 2004; Timaeus and Jasseh, 2004; Porter and Žaba, 2004).

Mortality conditions play a key role in future population composition thus the mortality component of projections for countries with generalized epidemics should include the influence of the future trajectory of the epidemic. Using tools like the United Nations EPP model (Brown et al., 2010; Ghys and Garnett, 2010) one can estimate or project features of an HIV epidemic such as prevalence and coverage of anti-retroviral therapy (ART) among the infected population, which can then be mapped onto a set of age-specific mortality rates. In this paper, we develop a model of age-specific mortality as a function of HIV prevalence and one of three possible inputs (life expectancy at birth, child mortality alone,  ${}_5q_0$ , or child mortality with adult mortality,  ${}_{45}q_{15}$ ) for the 39 countries of the world experiencing a generalized HIV epidemic.

## Data

We calibrate this model with a set of 312 5-year life tables for each sex from the World Population Prospects 2010 revision (United Nations, Department of Economic and Social Affairs, Population Division, 2011). This dataset contains mortality information from 39 countries with generalized HIV epidemics with eight 5-year life tables from 1970-2010 for each country.<sup>1</sup> All life tables have uniform 5-year age intervals except for the youngest age groups up to an open interval of 100 (0, 1-4, 5-9, 10-14,...,100+). We model the logged  ${}_n m_x$  column of these life tables. Finally, we obtain the mid-period (1973, 1978, 1983,...,2008) HIV prevalence as well as adult and child ART coverage for each of the 312 country-periods from UNAIDS estimates.

## Model

Our objective is a parsimonious model that can represent the age-pattern of mortality rates for countries with generalized HIV epidemics as a function of HIV prevalence and some other mortality indicator ( $e_0$ ,  ${}_5q_0$  alone, or  ${}_5q_0$  with  ${}_{45}q_{15}$ ). The general form of the model

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<sup>1</sup>The 39 countries were designated as countries with a generalized HIV epidemic if the country had  $> 1\%$  prevalence during any of the 5-year periods from 1970-2010.

represents a set of age-specific mortality rates as the weighted sum of three independent, age-varying components that represent the age-varying nature of the mortality schedule. A step-by-step discussion of the modeling strategy is below.

We first derive the age-varying components from a Singular Value Decomposition (SVD) of the matrix of observed mortality rate schedules. The resulting ‘left-singular vectors’ are the independent components we need, and they have the convenient property of encoding the bulk of the variation among the observed mortality schedules in a small number of these vectors.<sup>2</sup>

1. SVD components of life tables  $L$ :  $X_{i,x}$

We next regress (OLS) each of the 312 mortality rate schedules on the first few left-singular vectors from the SVD yielding a set of weights ( $\omega_{\ell,i}$ ) for each country-period life table that relate the individual mortality rate schedule to the components.

2. Project logged mortality rates onto SVD components  $X_{i,x}$ :

$$\lambda_{\ell,x} = c_{\ell} + \sum_{i=1}^3 \omega_{\ell,i} \cdot X_{i,x} \quad (1)$$

where

- $\lambda_{\ell,x}$  are logged age-specific mortality rates for life table  $\ell$ :  $\lambda_{\ell,x} = \log({}_n m_x)_{\ell}$
- $X$  are age-specific SVD components
- $\omega$  are weights specific to SVD components
- $c$  is a constant
- $\ell$  indexes life tables:  $\ell \in \{1, \dots, L\}$
- $i$  indexes age-specific SVD components:  $i \in \{1, 2, 3\}$
- $x$  indexes age:  $x \in \{0, 1, 5, 10, \dots, 100\}$

Once we have obtained a set of  $\omega$ ’s for each country-period life table, we model each  $\omega_i$  as a combination of HIV prevalence and life expectancy at birth or child mortality alone or child mortality with adult mortality. We use Bayesian Model Averaging (Raftery, 1995) to adjudicate amongst the various combinations of these variables to find the combination that best predicts each  $\omega_i$ .<sup>3</sup>

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<sup>2</sup>We tested the model using a varying number of components and found little improvement in fit when using more than three components. The lack of improvement with inclusion of higher order components likely reflects the fact that, similar to a Principal Components Analysis, each successive component accounts for successively smaller proportions of the overall variance.

<sup>3</sup>This step is performed with the function `bicreg()` from the package `BMA` (Raftery et al., 2012) in the statistical analysis software `R`. In model selection we included antiretroviral therapy coverage for both adults and children among the set of potential predictor variables but neither was found to be predictive of any of the weights.

### 3. Model $\omega$ 's :

$$\omega_{\ell,i} = \beta_{0,i} + \beta_{1,i} \cdot P_{\ell} + \beta_{2,i} \cdot e_{0,\ell} + \epsilon_{\ell,i} \quad (2)$$

or

$$\omega_{\ell,i} = \beta_{0,i} + \beta_{1,i} \cdot P_{\ell} + \beta_{2,i} \cdot {}_5q_{0,\ell} + \epsilon_{\ell,i} \quad (3)$$

or

$$\omega_{\ell,i} = \beta_{0,i} + \beta_{1,i} \cdot P_{\ell} + \beta_{2,i} \cdot {}_5q_{0,\ell} + \beta_{3,i} \cdot {}_{45}q_{15,\ell} + \epsilon_{\ell,i} \quad (4)$$

where

- $P$  is HIV prevalence
- $e_0$  is the expectation of life at birth
- ${}_5q_0$  is the probability of death from birth to 5
- ${}_{45}q_{15}$  is the probability of death from age 15 to 60

Equations 2-4 present models predicting a  $\omega$  as a function of all predictor variables, but BMA actually selects more parsimonious models. We also fit separate models for each  $\omega_i$  by sex as well as for African countries and non-African countries (Bahamas, Belize, Guyana, Haiti, Jamaica, Thailand). The region-sex-specific models for each  $\omega_i$  for the models based on life expectancy are presented below. For the  $e_0$  model, the same set of predictor variables is found through BMA for both regions and sexes. The region-sex-specific models and coefficients for all input combinations can be found in appendix 1.

$$\omega_{1,s,r} = \beta_{0,1,s,r} + \beta_{1,1,s,r} \cdot e_{0,s,r} \quad (5)$$

$$\omega_{2,s,r} = \beta_{0,2,s,r} + \beta_{1,2,s,r} \cdot e_{0,s,r} + \beta_{2,2,s,r} \cdot P_r \quad (6)$$

$$\omega_{3,s,r} = \beta_{0,3,s,r} + \beta_{1,3,s,r} \cdot e_{0,s,r} + \beta_{2,3,s,r} \cdot P_r \quad (7)$$

where

- $s$  indexes sex
- $r$  indexes region

Using equations 5 through 7 (or the models presented in appendix 1 for other inputs) we can predict  $\hat{\omega}$ 's, which when substituted in equation 1 produce a set of predicted mortality rates. The effective parameters in this model are the weights ( $\hat{\omega}$ ) as the components are age-specific and fixed.

4. Use models from step 3 to predict  $\hat{\omega}$ 's both in and out of sample.
5. Use predicted  $\hat{\omega}$ 's to predict  $\hat{\lambda}$ 's:

$$\hat{\lambda}_{\ell,x} = c_m + \sum_{i=1}^3 \hat{\omega}_{\ell,i} \cdot X_{i,x} \quad (8)$$

where

- $c_m$  is the median of the distribution of  $c_\ell$  values
- The ‘hat’ symbol over a variable, as in  $\hat{\omega}$ , indicates *predicted* value

## Model Validation

To assess the model’s ability to replicate mortality rate schedules for populations with generalized HIV epidemics, we use the model outlined above using all three input combinations to predict a set of age-specific mortality rates for each of the in-sample country-period life tables. We then compare the predicted mortality rates to those in the observed dataset. For comparison with existing models of all-age mortality, we fit the same set of tables using the WHO modified logit system (Murray et al., 2003), the UN model life tables for developing countries (United Nations, 1982) and the Coale-Demeny regional model life tables (Coale and Demeny, 1966; Coale and Guo, 1989). Table 1 presents the mean absolute error amongst all age groups and all country-periods along with the mean absolute error for life expectancy at birth ( $e_0$ ), under-5 mortality ( ${}_5q_0$ ), and adult mortality ( ${}_{45}q_{15}$ ). Because the system described in this paper is essentially an HIV-calibrated model life table system we refer to it as “HIV MLT” in the table below. For all metrics, smaller numbers are better and the smallest number in each column is bolded.

Table 1: Mean absolute error for three mortality indicators and amongst all-ages and life tables after fitting WPP life tables with various all-age mortality models.<sup>†</sup> The smallest error in each column is bolded.

Model	Male				Female			
	All-ages <sup>‡</sup>	$e_0$	${}_5q_0$	${}_{45}q_{15}$	All-ages <sup>‡</sup>	$e_0$	${}_5q_0$	${}_{45}q_{15}$
HIV MLT								
input: $e_0$	2.88	<b>0.88</b>	12.98	20.79	2.25	<b>0.90</b>	11.77	19.64
input: ${}_5q_0$	2.99	1.63	10.98	30.72	2.35	1.85	10.42	31.17
input: ${}_5q_0$ and ${}_{45}q_{15}$	<b>2.67</b>	1.08	<b>10.70</b>	<b>11.92</b>	<b>2.06</b>	1.22	<b>10.33</b>	<b>14.97</b>
WHO <sup>§</sup>	5.64	1.18	—	—	4.17	0.92	—	—
CD <sup>§</sup>	3.54	—	15.43	36.07	2.78	—	16.52	32.67
UN	5.99	1.18	12.47	37.67	11.03	1.43	26.54	36.86

<sup>†</sup> The mean absolute errors for ‘All-ages’,  ${}_5q_0$ , and  ${}_{45}q_{15}$  are expressed per 1,000

<sup>‡</sup> ‘All-ages’ refers to the mean absolute error for the non-logged mortality rates across age groups (0, 1-4, 5-9, 10-14, ..., 75) and amongst all life tables (1, ...,  $L = 312$ ).

<sup>§</sup> ‘WHO’ and ‘CD’ contain blank spaces as these quantities are inputs to these systems and thus have no error.

One can see from table 1 that our model is able to outperform these three existing model life table systems. This result is no surprise not because the method used to generate those

models is flawed but rather because the data sets used to calibrate those models do not contain data from countries and periods where the HIV epidemic was generalized. We show smaller errors on every metric for both sexes. The ‘HIV MLT’ model shows mean absolute errors for life expectancy of less than one year for both sexes with life expectancy at birth as an input along with slight errors in predicting the probability of childhood death of 0.014 and 0.012 for the male and female models respectively. The error in predicting adult death is slightly higher than for childhood death at about 0.02 for both sexes.

An advantage of this model and what tends to explain the overall better fit is its ability to produce the accentuated adult mortality humps associated with a generalized HIV epidemic that other models often miss. Figure 1 plots the predicted  ${}_n m_x$  schedule from the ‘HIV MLT’ model along with the WHO, Coale and Demeny, and UN fits for Lesotho females 2005-10. This figure makes clear that when HIV prevalence is high and the HIV hump is present, the ‘HIV MLT’ model is able to closely estimate age-specific mortality. Additional selected fits can be found in Appendix 2.

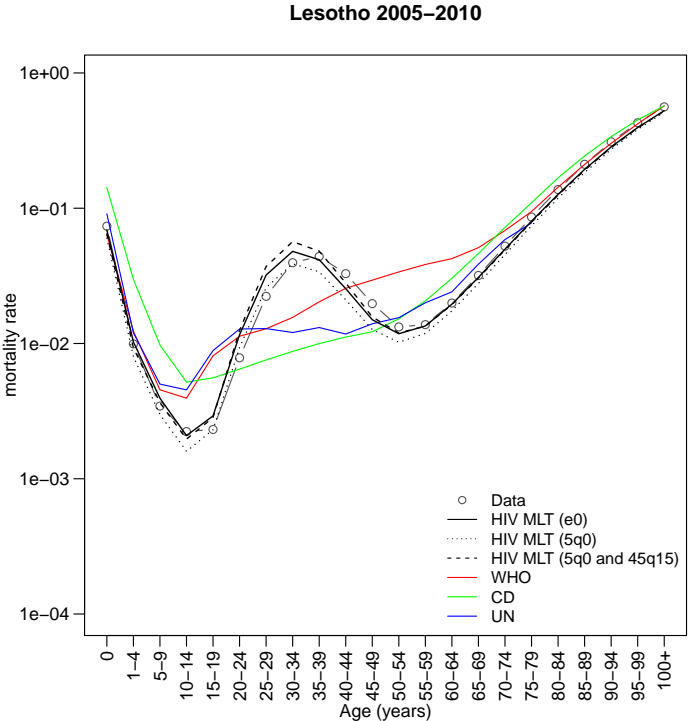


Figure 1: Fits of ‘HIV MLT’ model and other model life table systems for Lesotho females 2005-2010.

## Discussion

We present a flexible, parsimonious model of age-specific mortality for countries with generalized HIV epidemics. First, a set of age-specific mortality rates is represented as the weighted combination of a set of age-varying components. On a second level, the weights are modeled as a function of mortality indicators and epidemic characteristics. This structure allows us to map HIV prevalence combined with some other mortality indicator(s) onto a set of mortality rates reflecting the mortality impact of a generalized HIV epidemic.

We perform an in-sample validation method by predicting age-specific mortality for the 39 generalized HIV epidemic countries with the model presented here. Results suggest a three-component model fits best with modest errors for several mortality indicators.

Combined with the outputs of existing epidemiological and demographic models, this model makes it possible to estimate future mortality profiles for countries with generalized HIV epidemics.

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## Appendix 1: $\omega$ model coefficients

### HIV prevalence and life expectancy

Table 2: Coefficients for modeling weights ( $\omega_{i,s,r}$ ) as a function of  $e_0$  and prevalence

	Intercept	$e_0$	Prev
Female: Africa			
$\omega_{1,f,a}$	14.070	0.222	—
$\omega_{2,f,a}$	-3.511	0.051	0.180
$\omega_{3,f,a}$	2.491	-0.048	0.069
Female: Caribbean			
$\omega_{1,f,c}$	7.361	0.346	—
$\omega_{2,f,c}$	-6.110	0.093	0.493
$\omega_{3,f,c}$	2.962	-0.065	0.259
Male: Africa			
$\omega_{1,m,a}$	14.308	0.230	—
$\omega_{2,m,a}$	-3.461	0.053	0.173
$\omega_{3,m,a}$	2.195	-0.045	0.075
Male: Caribbean			
$\omega_{1,m,c}$	5.784	0.396	—
$\omega_{2,m,c}$	-5.968	0.098	0.517
$\omega_{3,m,c}$	2.316	-0.059	0.230

## HIV prevalence with child mortality alone

Table 3: Coefficients for modeling weights ( $\omega_{i,s,r}$ ) as a function of  ${}_5q_0$  and prevalence

	Intercept	${}_5q_0$	Prev
Female: Africa			
$\omega_{1,f,a}$	29.620	-24.009	-0.082
$\omega_{2,f,a}$	0.442	-7.653	0.148
$\omega_{3,f,a}$	-0.813	4.766	0.089
Female: Caribbean			
$\omega_{1,f,c}$	34.308	-54.315	—
$\omega_{2,f,c}$	1.258	-15.337	0.455
$\omega_{3,f,c}$	-2.227	11.282	0.293
Male: Africa			
$\omega_{1,m,a}$	29.957	-23.370	-0.086
$\omega_{2,m,a}$	0.522	-7.309	0.147
$\omega_{3,m,a}$	-0.891	4.697	0.090
Male: Caribbean			
$\omega_{1,m,c}$	34.883	-53.147	—
$\omega_{2,m,c}$	1.393	-14.712	0.461
$\omega_{3,m,c}$	-2.216	9.559	0.273

## HIV prevalence, child mortality, and adult mortality

Table 4: Coefficients for modeling weights ( $\omega_{i,s,r}$ ) as a function of  ${}_5q_0$ ,  ${}_{45}q_{15}$ , and prevalence

	Intercept	${}_5q_0$	${}_{45}q_{15}$	Prev
Female: Africa				
$\omega_{1,f,a}$	31.309	-16.548	-8.934	—
$\omega_{2,f,a}$	-0.201	-11.627	4.106	0.098
$\omega_{3,f,a}$	-1.300	1.757	3.109	0.051
Female: Caribbean				
$\omega_{1,f,c}$	37.358	-0.113	-36.203	—
$\omega_{2,f,c}$	0.823	-25.844	6.619	0.388
$\omega_{3,f,c}$	-2.572	2.974	5.234	0.241
Male: Africa				
$\omega_{1,m,a}$	32.128	-14.473	-9.935	—
$\omega_{2,m,a}$	-0.470	-12.233	5.040	0.094
$\omega_{3,m,a}$	-1.131	3.509	1.215	0.077
Male: Caribbean				
$\omega_{1,m,c}$	38.136	-26.672	-21.605	—
$\omega_{2,m,c}$	0.573	-22.347	5.959	0.419
$\omega_{3,m,c}$	-1.934	12.191	-2.054	0.287

## Appendix 2: Selected Fits

The following series of plots show example fits with this model. In this section we show a mix of regions, sexes, and prevalence levels.

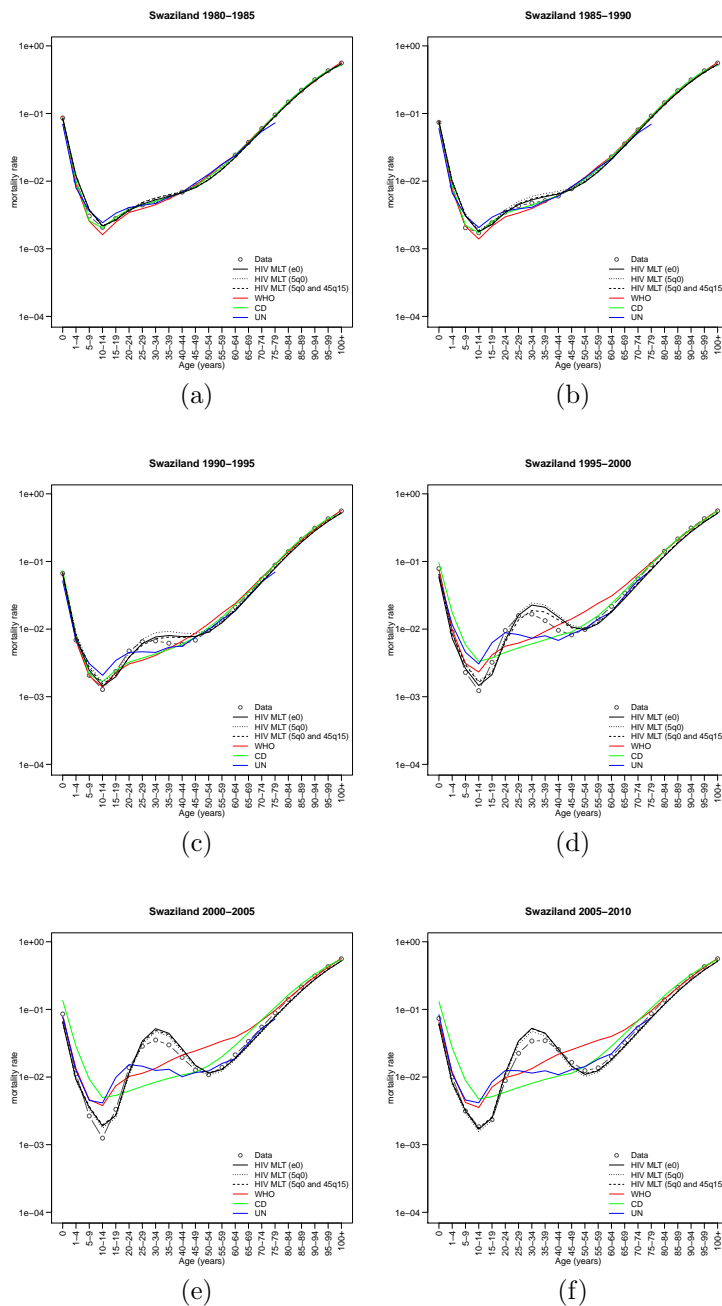


Figure 2: Swaziland, Females 1980-2010

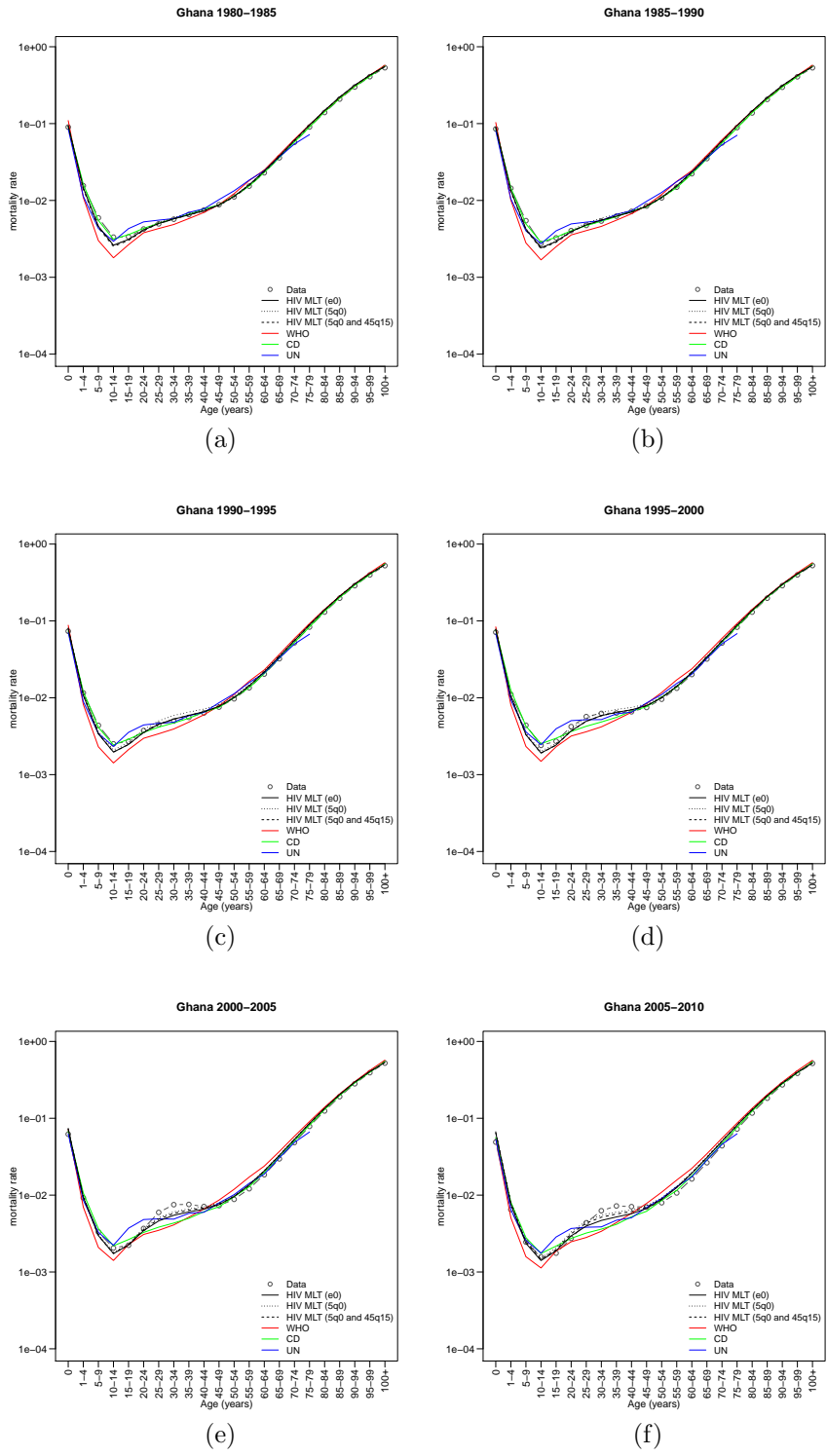


Figure 3: Ghana, Females 1980-2010

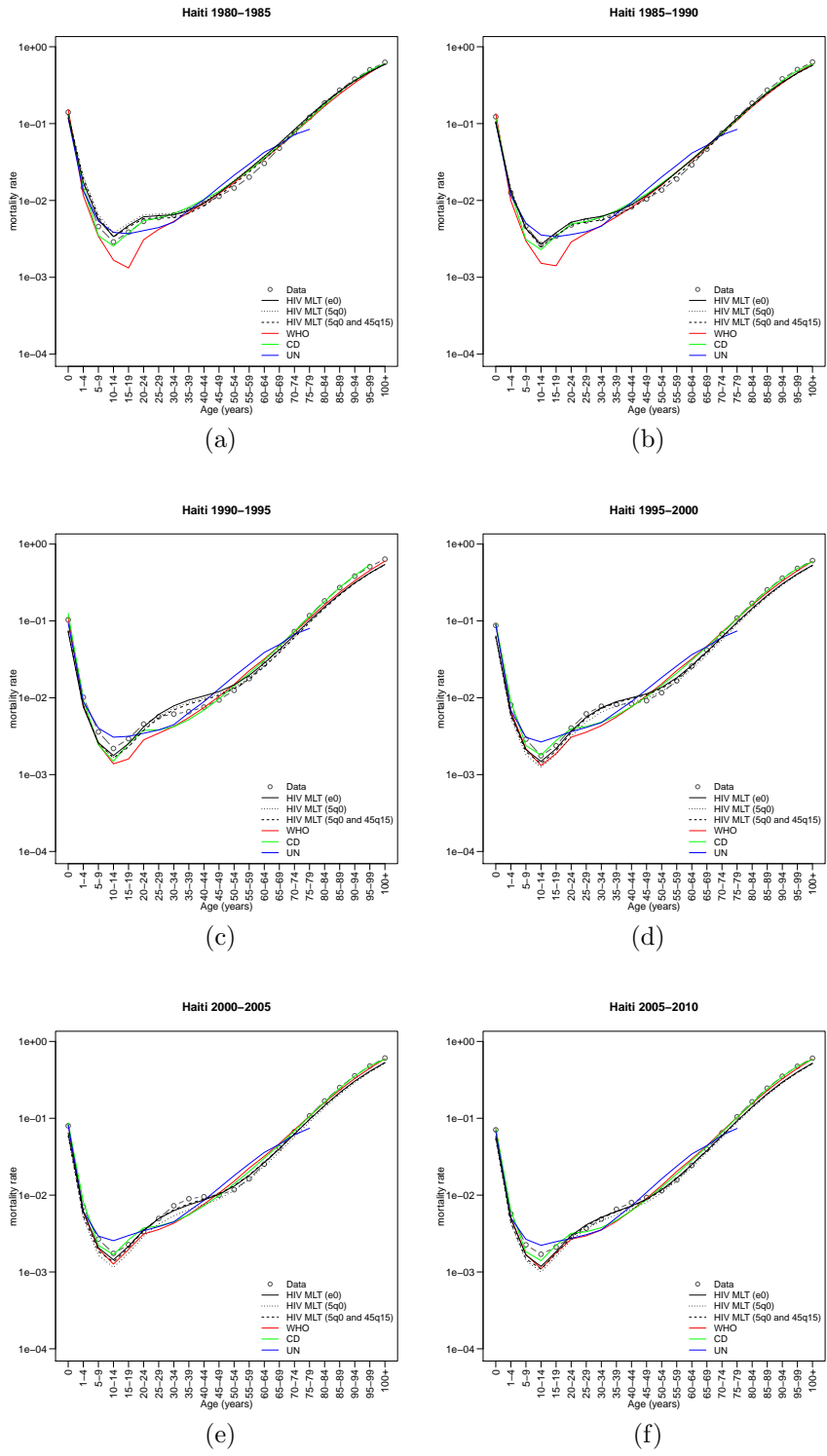


Figure 4: Haiti, Males 1980-2010

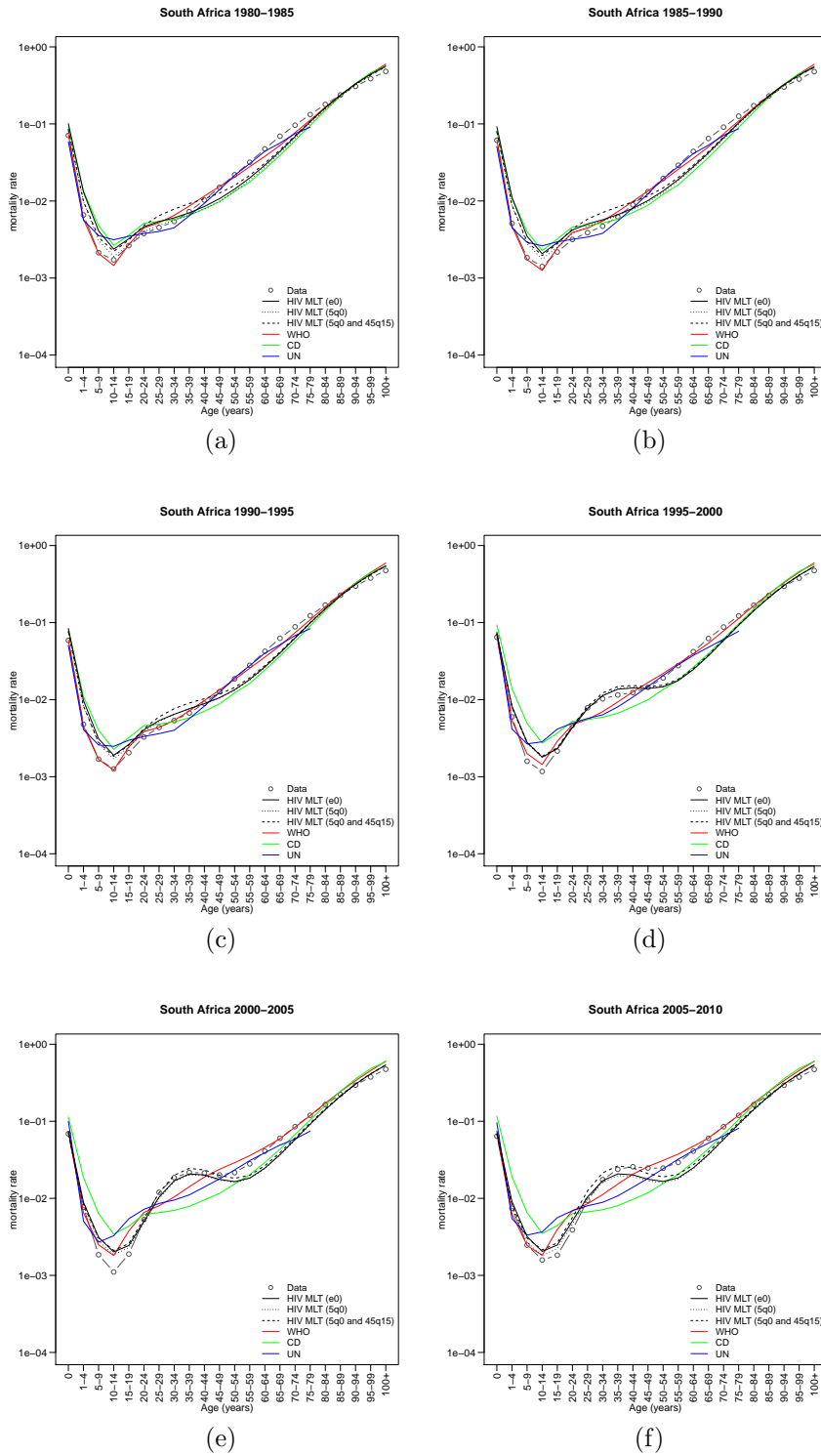


Figure 5: South Africa, Males 1980-2010