

# **Influenza mortality in the United States, 2009: Burden and timing**

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## **150-word abstract**

In April 2009, the most recent pandemic of influenza A began. We present the first estimates of pandemic mortality based on the newly-released final data on deaths in 2009. In the aggregate, 2009 was not an unusual year for pneumonia and influenza mortality. Compared to the typical pattern of seasonal flu deaths, pneumonia and influenza age-specific mortality in 2009, as well as influenza-attributable (excess) mortality, skewed much younger. In many age groups, pneumonia and influenza mortality in October and November 2009 broke month-specific records since 1959 when the current series of detail US mortality data began. Mortality in influenza pandemics skews younger than seasonal influenza. This can be explained in part by a protective effect due to antigenic cycling. If older cohorts have been previously exposed to a similar antigen, immune memory can result in lower death rates at older ages.

## **Abstract**

### **Background**

In April 2009, the most recent pandemic of influenza A began. We present the first estimates of pandemic mortality based on the newly-released final data on deaths in 2009 and 2010 in the United States.

### **Methods**

We obtained data on influenza and pneumonia deaths from the National Center for Health Statistics (NCHS). Age- and sex-specific death rates, and age-standardized death rates, were calculated. Using standard Serfling-type methods, excess mortality was calculated separately by age groups.

### **Results**

In the aggregate, 2009 was not an unusual year for pneumonia and influenza mortality. Compared to the typical pattern of seasonal flu deaths, pneumonia and influenza age-specific mortality in 2009, as well as influenza-attributable (excess) mortality, skewed much younger. In many age groups, pneumonia and influenza mortality rates in October and November 2009 broke month-specific records since 1959 when the current series of detail US mortality data began.

## Conclusions

Mortality in influenza pandemics skews younger than seasonal influenza. This can be explained in part by a protective effect due to antigenic cycling. When older cohorts have been previously exposed to a similar antigen, immune memory results in lower death rates at older ages. Age-targeted vaccination should be considered in future pandemics.

## Introduction

In April 2009, a novel strain of influenza A/H1N1 emerged in Mexico and rapidly spread to the United States [1] and then worldwide. On 11 June 2009, the World Health Organization declared a pandemic [2]. The pandemic would eventually cause an estimated 284,500 deaths, worldwide [3]. Approximately 20% of the US population contracted influenza during the pandemic [4,5]. Overall, the pandemic case fatality rate was low, and attack rates were higher among children and young adults [6]. The sparing of adults 65 and older is thought to be due to immunity from previous exposure to antigenically-similar H1N1 strains [7–12].

In this article, we analyze final mortality data for the United States which were released in August 2012. This is the first analysis of the complete mortality record of influenza for 2009. Our age- and sex-specific analysis incorporates data from 1959–2010. Pneumonia and influenza deaths are commonly analyzed together [13], and in 2009, there were 53,692 pneumonia and influenza deaths, making this combination the eighth leading cause of death [14]. Age-specific mortality burden estimates, like those presented here, can aid pandemic planning.

## Results

Figure 1 plots the age-standardized death rate (ASDR) for pneumonia and influenza for the United States from January 1959 to December 2010. This figure presents the most complete monthly record of pneumonia and influenza ASDR in the US. Considering death rates weighted over all ages, recent influenza pandemics are not especially severe. This is seen clearly in figure 1 for 2009, as well as for the H3N2 “Hong Kong” influenza pandemic of 1968–69. Due to the fall wave of the 2009 H1N1 pandemic, the mortality of the winter 2009–10 influenza season began unusually early [15], creating an atypical wide-and-broad pattern. Calendar year 2003 is the only other recent year with pneumonia and influenza mortality rising so strongly in the fall (in this case, December), which was associated with the emergence of the Fujian strain of influenza A/H3N2 [16]. Other remarkable features of figure 1 are the secular decline of the ASDR for pneumonia and influenza, and the consistency with which male death rates exceed those of females.

Figure 2 presents the distribution of pneumonia and influenza deaths by age group (0, 1–4, 5–9, . . . , 95–99,  $\geq 100$ ) and by sex. The solid bars represent calendar year 2009,

and, for comparison, the cross-hatched bars are for the average of 1999–2008. Mortality skewed younger in 2009 compared to the previous 10-year average: the solid bars exceed the hatched bars for all ages 0–69 (to 74 for females). The female histograms skew older than those for males. Although females consistently have lower pneumonia and influenza death rates than males, their longer life expectancy [14] means more females at older ages, and therefore more deaths at older ages.

Figure 3 plots pneumonia and influenza death rates versus age for each month of 2009, separately by sex. This illustrates combined aspects of figure 1 (unusual timing) and figure 2 (younger age distribution). As table 1 details, October–November 2009 is one of the most remarkable periods in the history of pneumonia and influenza mortality since 1959, especially for females. Out of all Octobers, 1959–2010, October 2009 recorded the highest pneumonia and influenza death rate for females in the age span 25–59. This month was less notable for males, although in ages 25–29 and 35–39 they also had a record mortality rate. Likewise, November 2009 recorded record-high pneumonia and influenza death rates: for females, in the age spans 20–34 and 40–59, and for males, 25–29. December 2009, on the other hand, did not experience unusually high mortality rates (figure 1). For Decembers, of the 44 age groups (22 age groups  $\times$  2 sexes), 31 have their record high in 1968, coinciding with the H3N2 “Hong Kong” pandemic [17]. Note that these comparisons refer to mortality rates, not death counts; the latter rise with population growth. Prior to 1959, age  $\times$  sex  $\times$  cause-of-death monthly data are unavailable, so the October–November record-high death rates are at least 52-year records.

Figure 4 displays the age distribution of influenza-attributable excess mortality as calculated by a Serfling-type regression model of pneumonia and influenza death rates. The age groups and shading follow that of figure 2. Even more so than age-specific mortality (i.e., figure 2), 2009 pandemic excess mortality skews young. Thus, the youngness of the distributions in 2009 (solid) versus 1999–2008 (cross-hatched) reflects the pandemic, and is not an artifact of switching the metric from raw mortality (figure 2) to excess mortality (figure 4).

Table 2 provides calendar year 2009 excess pneumonia and influenza mortality estimates by age group and sex, with the corresponding 95% confidence bounds. Negative excess mortality values indicate that the observed pneumonia and influenza deaths in the age group were less than the age group-specific Serfling baseline. Where the lower bound is negative, it may be interpreted as non-statistically-significant excess mortality. Excess mortality is highest in the age group 55–59 for males and 50–54 for females. For both sexes, there is no significant excess mortality at age 80 and above. Below age 80, two age groups for males (1–4, 60–64), and three for females (0, 70–74, 75–79), were not significant. The ratio of excess mortality to all pneumonia and influenza mortality is comparable to similar ratios in prior work [18].

## Discussion

Overall, 2009 had about 2,010 pneumonia and influenza excess deaths. It was not an unusually lethal flu year; the confidence bounds on the total excess deaths include zero (cf. table 2). This should not be taken to mean that the pandemic was not deadly, especially below age 60. The bulk of pneumonia and influenza mortality occurs above age 60, and when the pandemic strain of influenza H1N1 displaced other circulating strains [19], the accompanying rise in pneumonia and influenza mortality took place in age groups with generally low mortality rates. Indeed, the unusual mortality at young ages is a remarkable indication of the specific impact of the pandemic. Moreover, this number should not be regarded as contradicting other estimates of mortality burden [4] since multiplier models (see also [20]) and Serfling-type models represent different approaches.

The public health concern about influenza pandemics stems, above all else, from their potential to kill millions, as in the 1918–19 pandemic [21]. The mortality of the 2009 pandemic was unusual not for its severity but for its timing and age distribution. In terms of the ASDR and overall excess mortality, the 2009 pandemic was not distinguishable from a typical influenza season. The H3N2 (“Hong Kong”) pandemic of 1968–69 was, similarly, comparable in ASDR to the previous winter (figure 1). However, in October and November 2009, pneumonia and influenza mortality rates in a number of age groups were the highest for those months since at least 1959. This phenomenon, which occurred exclusively below age 60, was more pronounced for females; the reasons for this are unclear.

Seasonal influenza typically kills with an old-age mortality pattern: the cross-hatched bars (1999–2008 average) of figures 2 and 4 peak at age 80 and above. There were three influenza pandemics in the twentieth century (1918–19, 1957, 1968–69) [22], all of which skewed young [23–28]. The 2009 data confirm this pattern. There is a limited sample from which to draw conclusions, but influenza pandemic mortality consistently skews young. The younger pattern is seen in both 2009 excess mortality (figure 4 and table 2) and age-specific mortality (figure 2, figure 3, and table 1). Our finding is also consistent with earlier work [29]. Excess mortality is calculated relative to a baseline model [30,31]. Thus, excess mortality estimates may be influenced by timing (phase) as well as severity (amplitude). The age-specific mortality data are not model-based, so the qualitative agreement of younger skew in both modes of analysis suggests that the skew is not an artifact.

After reemergence in 1977, influenza A/H1N1 spread in children and young adults because of its remarkable antigenic similarity to the pre-1957 H1N1 strains [32,33]. The protective effect of older cohorts having seen this strain before was clear-cut. In the case of 2009, an analogous phenomenon is likely [7,8]. The relative peak at age 25–29 in 2009 (figure 4) may be an example of original antigenic sin [34–39]. Specifically, persons who were first exposed to influenza during the H1N1-dominant flu seasons of 1978–79, 1983–84, 1986–87, or 1988–89 [40] may have reacted differently than surrounding cohorts. This is somewhat speculative and merits further investigation. The local maximum, for example,

is more pronounced for females, which is not accounted for by the original antigenic sin explanation. Our data do not record prior infection status, so more in-depth exploration would require additional data.

One of the biggest challenges of pandemic preparedness is rapid formulation and manufacture of a strain-specific vaccine [41]. Our analysis suggests that younger ages (30–60, see figure 4) should be prioritized in the event of a pandemic. Excess pneumonia and influenza mortality in 2009 peaked below age 60 for both sexes. The mortality data for 2009 do not provide the last word, but do suggest that age-targeted vaccination is a strategy well worth considering.

## Materials and Methods

We obtained data on number of deaths, by cause, from the mortality detail files of the National Center for Health Statistics (NCHS) [42]. Deaths were stratified by age, sex, month, and underlying cause. We extracted data on influenza and pneumonia deaths from January 1959 to December 2010. As shown in table 3, this period spans four revisions of the International Classification of Diseases (ICD 7–10). To ensure comparability, all data were converted to ICD-10 using the published crossover tables [43–45]. Prior to 1959, no detailed mortality data (i.e., simultaneously disaggregated by age, sex, month, and cause) are available. Final mortality data for 2009 and 2010 were released in 2012.

Rates (per 100,000) were calculated using these death counts in the numerator, and exposure data (person-years at risk) from the Human Mortality Database [46] in the denominator. The age-specific exposure data were interpolated to monthly units to match the death counts. Variable days per month (including leap years) were used in the monthly exposure interpolation.

Excess mortality was calculated using standard Serfling-type methods [30, 31] on the monthly death rates described above. Our excess mortality calculation was done separately by age groups, which permits full use of the data, while being collapsible to all-ages in a robust way [47]. We used a Serfling approach similar to that of [18], running regressions of the type:

$$\log \left( \frac{\text{deaths}_{a,s,t}}{\text{exposure}_{a,s,t}} \right) = \beta_0 + \beta_1 t + \beta_2 t^2 + \beta_3 t^3 + \beta_4 t^4 + \beta_5 t^5 + \beta_6 \cos \left( \frac{2\pi t}{12} \right) + \beta_7 \sin \left( \frac{2\pi t}{12} \right)$$

where subscripts  $a, s, t$  denote age- and sex-specific models, with monthly time ( $t$ ) resolution. The model de-trends by a quintic polynomial in time ( $t..t^5$ ), and includes two harmonic (sin, cos) terms;  $\beta_0.. \beta_7$  are coefficients to be estimated. Ordinary least squares was used; this technique produces comparable results to other approaches [48]. The models establish a baseline by estimating a waveform from the summer troughs (May–October) of pneumonia and influenza mortality. September and October 2009 were excluded from the baseline calculations due to the atypically early circulation of influenza virus during this time period.

Once the coefficients are estimated, a baseline is calculated from the regression equation, exponentiating to restore correct scale. Upper and lower bounds were calculated from the standard error of the prediction equation and the tails of the  $t$ -distribution (probability 0.025 in each tail), to create a 95% confidence bound. Excess mortality rates are the observed rates minus the baseline. Excess mortality is calculated as the excess mortality rate multiplied by the exposure. An analogous approach to excess mortality calculation is to use all-cause mortality in lieu of pneumonia and influenza (both as input to the baseline estimation, and as the rates from which the baseline is subtracted), but we chose to emphasize specificity over sensitivity [40]. Analysis was done using IDL version 8.2 (Exelis Visual Information Solutions, Inc., Boulder CO, USA) and Stata version 10.1 (StataCorp LP, College Station TX, USA).

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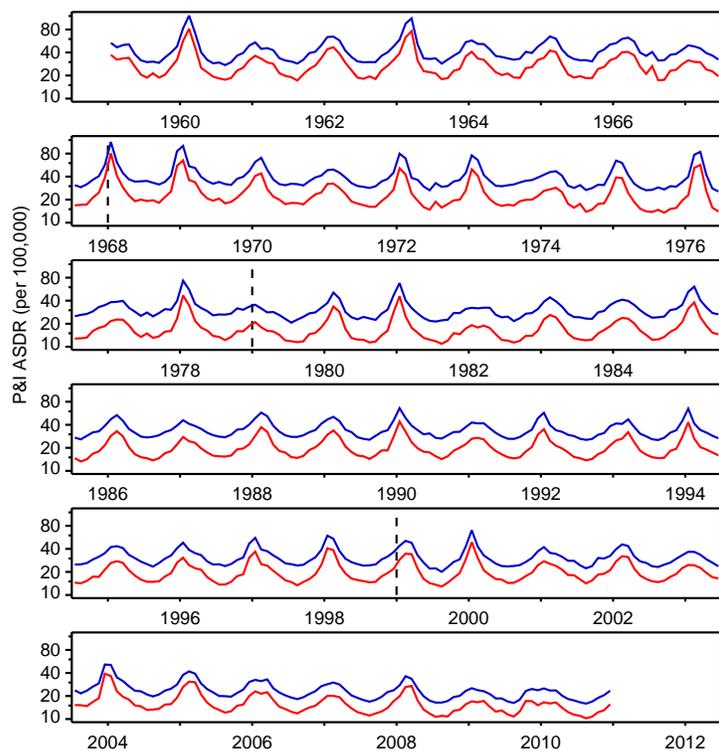
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## Figure Legends



**Figure 1. Monthly age-standardized death rate (ASDR) for pneumonia and influenza, United States, 1959–2010.** Males blue, females red. Dashed lines denote boundaries of ICD regimes, but data are adjusted for ICD changes.

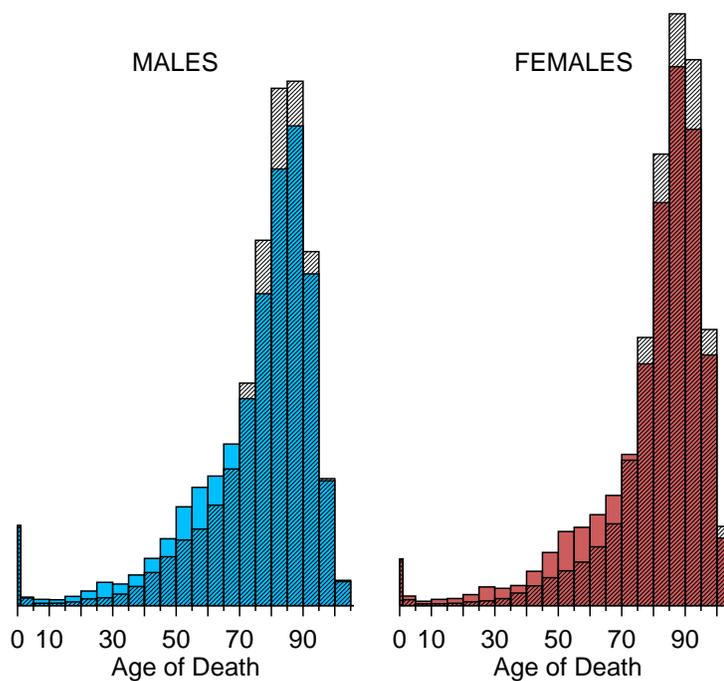
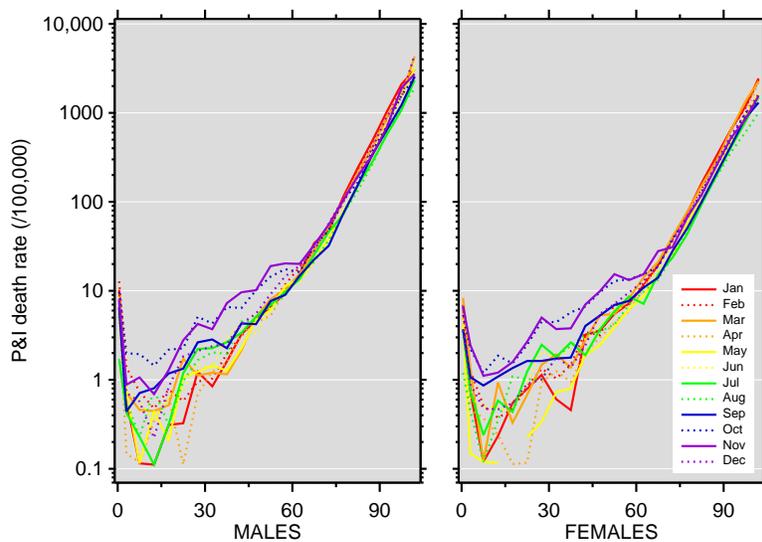


Figure 2. Histograms of deaths by age, pneumonia and influenza (ICD10 J10–J18), United States. 2009 is solid; 1999–2008 average, cross-hatched.



**Figure 3. Pneumonia and influenza (P&I) age-mortality profiles by month.** Vertical axis: P&I death rate per 100,000; horizontal axis: age. The gap in May for females reflects that there were no deaths in the 15–19 age group.

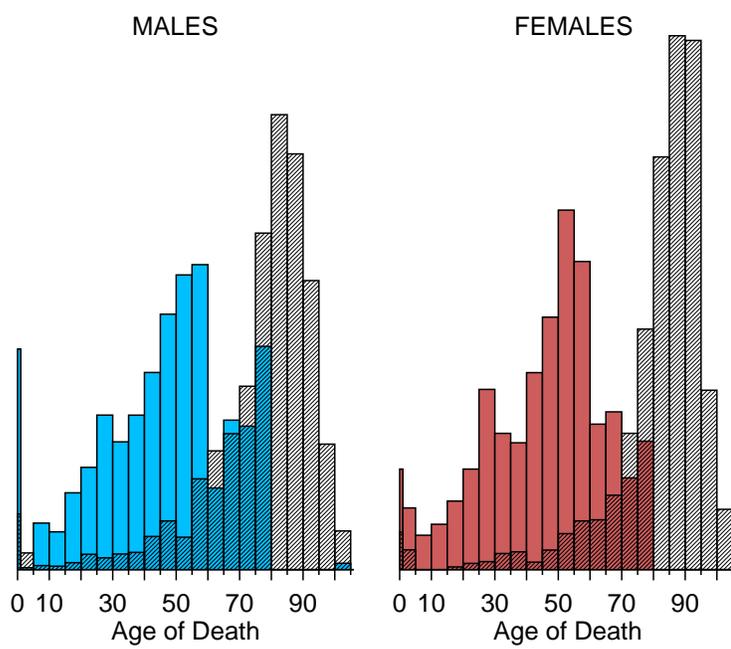


Figure 4. Histograms of excess deaths by age, pneumonia and influenza, United States. 2009 is solid; 1999–2008 average, cross-hatched.

## Tables

**Table 1. October and November record-high P&I mortality rates**

Age Group	Males				Females			
	Octobers		Novembers		Octobers		Novembers	
	Rate	Year	Rate	Year	Rate	Year	Rate	Year
0	175.98	1968	217.28	1959	145.57	1960	175.27	1968
1–4	12.56	1959	14.53	1960	12.18	1959	12.90	1960
5–9	2.44	1964	2.45	1969	2.78	1959	2.90	1964
10–14	1.53	1965	2.21	1962	1.88	1965	2.50	1962
15–19	2.27	1970	2.49	1959	2.19	1967	1.95	1969
20–24	2.60	1966	3.61	1961	3.20	1959	2.67	2009
25–29	5.11	2009	4.25	2009	4.51	2009	5.01	2009
30–34	4.42	1969	5.00	1965	4.44	2009	3.72	2009
35–39	6.58	2009	7.30	1968	5.73	2009	4.21	1959
40–44	9.05	1964	9.77	1968	6.67	2009	7.01	2009
45–49	13.69	1964	13.79	1969	9.69	2009	9.91	2009
50–54	18.23	1965	19.74	1968	12.91	2009	15.40	2009
55–59	24.71	1965	26.54	1968	13.49	2009	13.21	2009
60–64	37.55	1963	43.56	1968	21.00	1961	20.35	1960
65–69	62.33	1961	65.43	1964	28.07	1962	30.58	1959
70–74	108.52	1964	113.23	1968	50.47	1965	51.30	1965
75–79	206.63	1965	212.16	1965	111.15	1962	118.76	1961
80–84	371.10	1965	378.92	1972	226.37	1964	244.23	1961
85–89	780.87	1987	800.08	1991	458.96	1961	497.82	1961
90–94	1431.85	2001	1520.08	1995	891.89	1988	1032.25	1962
95–99	2588.80	1989	2926.11	2003	1763.49	1961	1755.16	1991
≥100	5638.09	2002	5232.87	2001	3380.34	2001	3406.07	1966

Record-high mortality rates (per 100,000) for 22 age groups in Octobers and Novembers, 1959–2010. Out of 88 age × sex × month combinations, 17 record highs occur in 2009, all below age 60.

**Table 2. Excess mortality estimates**

Age Group	Males		Females	
	Excess Mortality	(lower bound, upper bound)	Excess Mortality	(lower bound, upper bound)
0	38.8	(19.5, 55.2)	16.3	(-0.3, 30.2)
1-4	5.3	(-9.8, 17.2)	40.5	(27.7, 50.1)
5-9	39.5	(32.1, 44.8)	29.0	(22.6, 33.7)
10-14	32.3	(23.9, 38.4)	39.5	(29.2, 46.9)
15-19	61.9	(52.6, 68.6)	53.1	(44.7, 59.2)
20-24	81.6	(63.6, 95.0)	77.5	(64.2, 87.5)
25-29	126.0	(103.8, 143.8)	136.3	(116.4, 151.2)
30-34	106.1	(85.9, 122.7)	105.6	(85.5, 121.5)
35-39	129.0	(102.6, 151.5)	102.2	(80.8, 120.0)
40-44	165.2	(127.5, 198.2)	162.6	(121.5, 196.5)
45-49	214.2	(167.9, 255.7)	213.5	(161.5, 258.7)
50-54	258.3	(192.7, 317.7)	296.3	(225.8, 357.8)
55-59	254.1	(186.9, 316.3)	254.1	(185.9, 314.8)
60-64	70.9	(-11.6, 147.6)	133.8	(62.5, 199.4)
65-69	130.8	(44.7, 211.6)	120.3	(39.1, 195.6)
70-74	101.1	(-4.5, 200.8)	93.5	(-5.4, 186.2)
75-79	197.1	(48.2, 337.9)	117.7	(-30.9, 257.5)
80-84	-10.7	(-211.2, 180.2)	-8.6	(-244.3, 214.5)
85-89	-229.2	(-481.6, 10.0)	-242.4	(-581.8, 78.1)
90-94	-259.1	(-485.2, -47.6)	-210.9	(-531.6, 90.8)
95-99	-8.0	(-134.0, 105.4)	-290.2	(-544.2, -56.8)
≥100	2.9	(-52.5, 47.0)	-78.1	(-207.1, 32.4)
Total	1508.1	(-138.5, 3018)	1161.6	(-878.2, 3025.8)

Excess pneumonia and influenza mortality estimates (deaths) and 95% confidence bounds for calendar year 2009.

**Table 3. ICD codes for Pneumonia & Influenza**

Years	Causes used as Pneumonia & Influenza
1959–1967 (ICD 7)	480–483, 490–493
1968–1978 (ICD 8)	470–474 , 480–486
1979–1998 (ICD 9)	480–487
1999–2010 (ICD 10)	J10–J18

Codes from the four revisions of the ICD (International Classification of Diseases) merged to the combined cause “pneumonia and influenza.”